

Heavy Tails: Performance Models and Scheduling Disciplines

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based on joint ITC 18 Tutorial with Onno Boxma, Sem Borst and Mor Harchol-Balter

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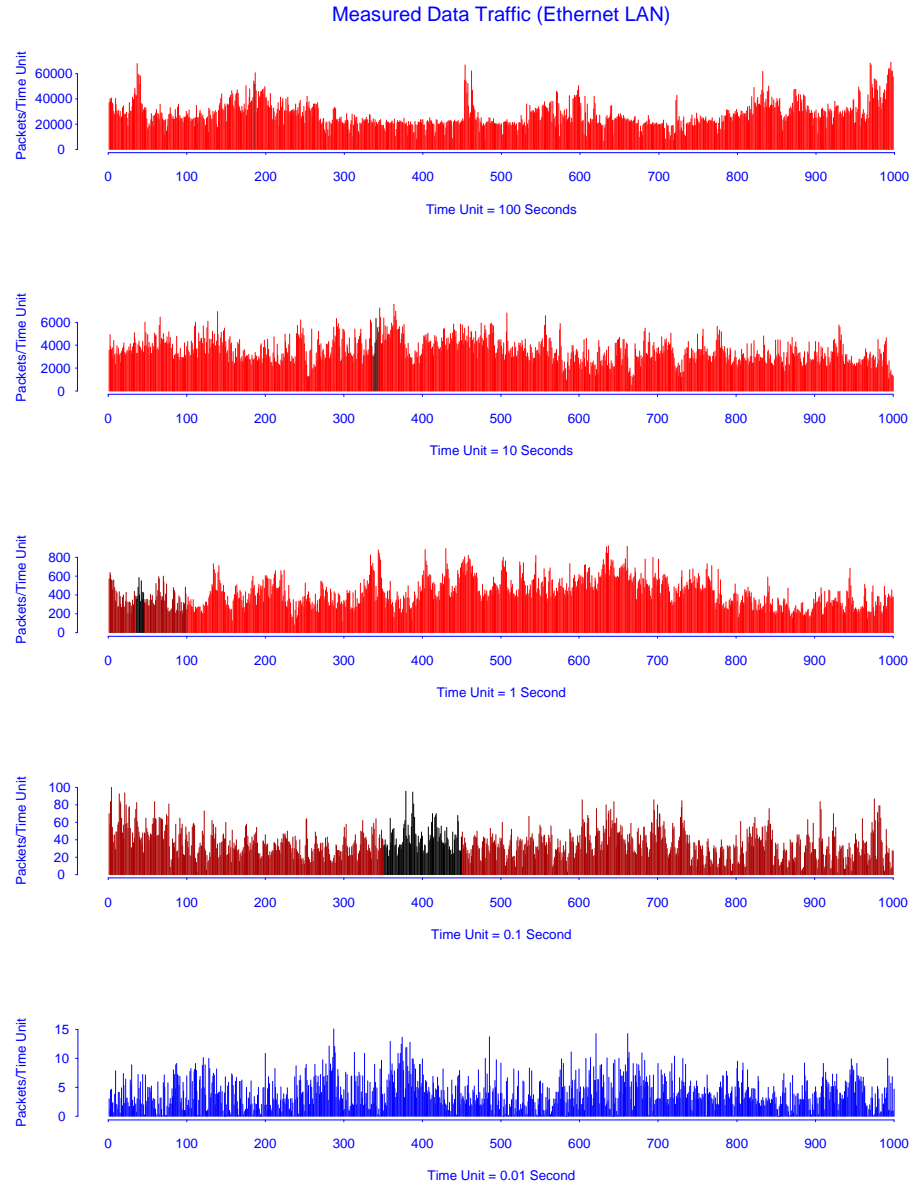
Part I – Introduction and Methodology

Heavy Tails: Performance Models and Scheduling Disciplines Part I – Introduction and Methodology

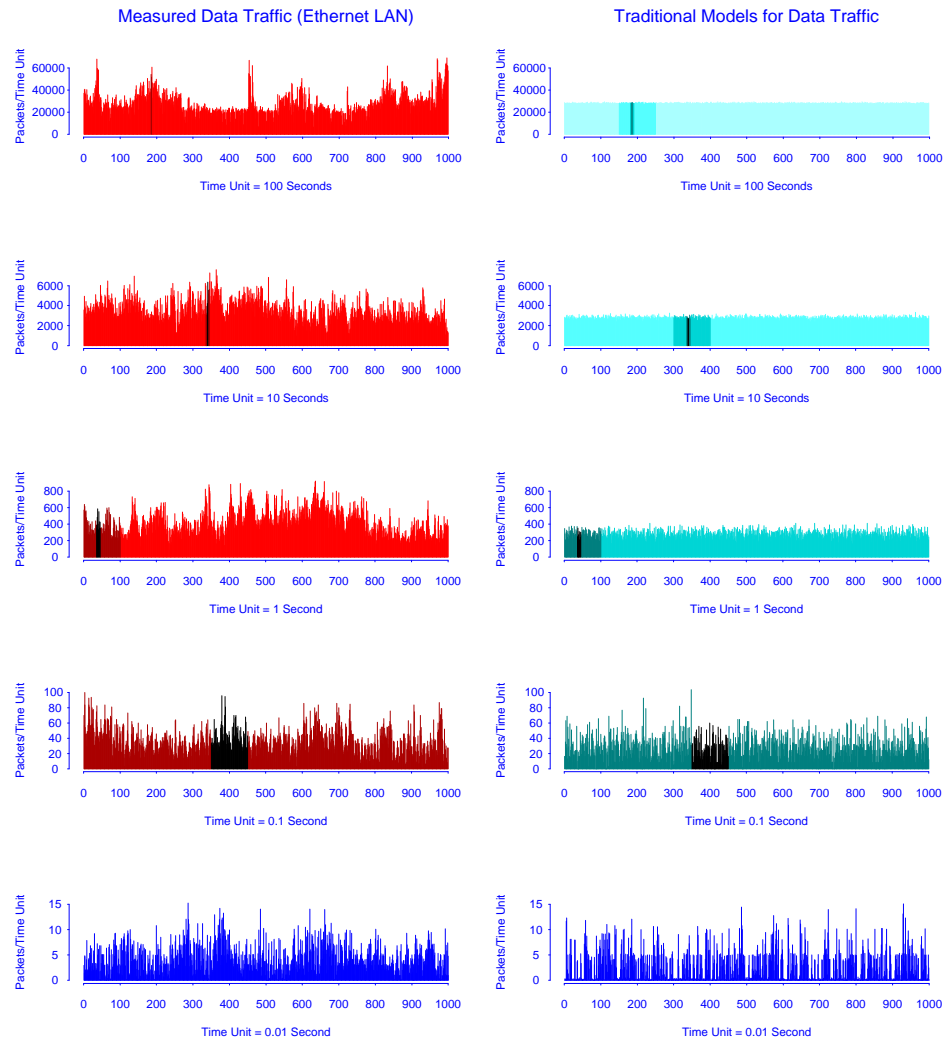
Tales to tell:

- traffic measurements and statistical analysis
- traffic modeling
- heavy-tailed input: performance analysis
- heavy-tailed input: damage control

traffic measurements and statistical analysis



traffic measurements and statistical analysis



The Bellcore data (courtesy: [W. Willinger](#))

Heavy Tails: Performance Models and Scheduling Disciplines Part I – Introduction and Methodology

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- traffic measurements and statistical analysis
- **traffic modeling**
- heavy-tailed input: performance analysis
- heavy-tailed input: damage control

traffic modeling

W is **heavy-tailed** if $e^{\epsilon x} P(W > x) \rightarrow \infty, \forall \epsilon > 0$.

Examples:

Pareto: $P(B > x) = \left(\frac{\theta}{\theta+x}\right)^\nu, \quad x \geq 0$.

Weibull: $P(B > x) = e^{-(x/b)^c}, \quad 0 < c < 1, \quad x \geq 0$.

lognormal: $P(B > x) = \int_x^\infty \frac{1}{y\sigma\sqrt{2\pi}} \exp\left[-\frac{[\log(y/m)]^2}{2\sigma^2}\right] dy, \quad x \geq 0$.

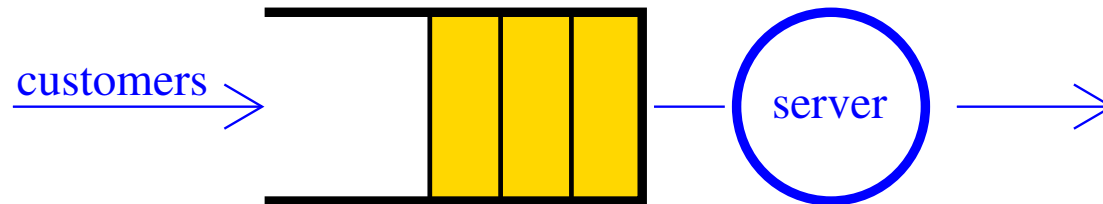
Heavy Tails: Performance Models and Scheduling Disciplines Part I – Introduction and Methodology

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heavy-tailed input: performance analysis

M/G/1



Key question: tail behavior of key performance measures

if

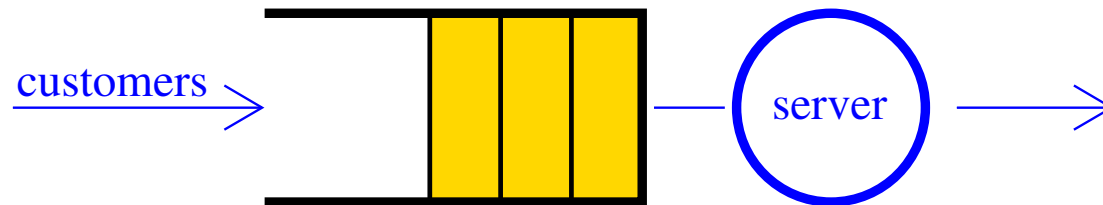
$$P(B > x) \sim x^{-\nu}, \quad x \rightarrow \infty,$$

then

$$P(W_{FCFS} > x) \sim ??, \quad x \rightarrow \infty.$$

heavy-tailed input: performance analysis

M/G/1



Key question: tail behavior of key performance measures

if

$$P(B > x) \sim x^{-\nu}, \quad x \rightarrow \infty,$$

then

$$P(W_{FCFS} > x) \sim Cx^{1-\nu}, \quad x \rightarrow \infty.$$

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heavy-tailed input: performance analysis/damage control

Key questions:

- tail behavior of key performance measures
- multiclass systems: effect of one class on another class

heavy-tailed input: performance analysis/damage control

Key questions:

- tail behavior of key performance measures
- multiclass systems: effect of one class on another class
- influence of the service discipline

heavy-tailed input: performance analysis/damage control

M/G/1

if

$$P(B > x) \sim x^{-\nu}, \quad x \rightarrow \infty,$$

then

$$P(W_{FCFS} > x) \sim Cx^{1-\nu}, \quad x \rightarrow \infty.$$

but perhaps

$$P(W_{smart} > x) \sim Dx^{-\nu}, \quad x \rightarrow \infty?$$

Ref.: V. Anantharam, QUESTA 33, 1999

heavy-tailed input: performance analysis/damage control

Key questions:

- tail behavior of key performance measures
- multiclass systems: effect of one class on another class
- influence of the service discipline
- fairness

Organization

- Part I. Methodology
- Part II. Workload asymptotics of GPS systems (sample path analysis)
- Part III. Delay asymptotics for various scheduling strategies; PS in integrated services environments

Part I

1. introduction

2. regular variation

3. $M/G/1$ FCFS: workload and waiting time

4. $M/G/1$: busy period

5. $M/G/1$ LCFS preemptive resume: sojourn time

6. epilogue

2. Regular variation

M/G/1 queue: $B(x) = P(B < x)$,

LST $E[e^{-sB}]$, mean $EB < \infty$.

arrival rate λ , load $\rho := \lambda EB < 1$.

Regularly varying service time distribution:

$$P(B > x) = x^{-\nu} L(x).$$

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Regularly varying service time distribution:

$$P(B > x) = x^{-\nu} L(x).$$

$L(\cdot)$ is slowly varying: $\lim_{t \rightarrow \infty} \frac{L(tx)}{L(t)} = 1, \forall x > 0$.

(we mainly consider $1 < \nu < 2$: $\text{Var}(B) = \infty$.)

Example: Pareto. $P(B > x) = \left(\frac{\theta}{\theta+x}\right)^\nu, \quad x \geq 0$.

Bingham & Doney

Key lemma for regularly varying tails

Let $n < \nu < n + 1$.

Equivalent are:

$$P(Y > x) \sim x^{-\nu} L(x), \quad x \rightarrow \infty,$$

and

$$E[e^{-sY}] - \sum_{j=0}^n EY^j \frac{(-s)^j}{j!} \sim -\Gamma(1 - \nu) s^\nu L\left(\frac{1}{s}\right), \quad s \downarrow 0.$$

3. $M/G/1$ FCFS: waiting time (and workload)

$G/G/1$ FCFS

Cohen (JAP, 1973):

$$P(B > x) \sim x^{-\nu} L(x), \quad x \rightarrow \infty, \quad \nu > 1 \iff$$

$$P(W > x) \sim \frac{\rho}{1-\rho} \frac{1}{\nu-1} \frac{1}{EB} x^{1-\nu} L(x), \quad x \rightarrow \infty.$$

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Cohen (JAP, 1973):

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$$P(W > x) \sim \frac{\rho}{1-\rho} \frac{1}{\nu-1} \frac{1}{EB} x^{1-\nu} L(x), \quad x \rightarrow \infty.$$

In fact (Pakes, JAP 1975, for subexponential B^{res}):

$$P(W > x) \sim \frac{\rho}{1-\rho} P(B^{res} > x), \quad x \rightarrow \infty.$$

Note that $P(B^{res} > x) = \int_x^\infty \frac{P(B > u)}{EB} du$

and $EB^{res} = \frac{EB^2}{2EB}$.

Four proofs for $M/G/1$

(a) Direct proof

(b) Proof via LST and Bingham-Doney

(c) Proof via sample-path argument

(d) Proof via conditional moments method (see Part III)

(a) Direct proof

$$P(W > x) \sim \frac{\rho}{1 - \rho} P(B^{res} > x), \quad x \rightarrow \infty.$$

See next slide:

$$W = {}^d B_1^{res} + \dots + B_K^{res},$$

with $K \sim \text{geom}(\rho)$.

Hence

$$\begin{aligned} P(W > x) &= (1 - \rho) \sum_{n=0}^{\infty} \rho^n P(B_1^{res} + \dots + B_n^{res} > x) \\ &\sim (1 - \rho) \sum_{n=0}^{\infty} \rho^n n P(B^{res} > x) \\ &= \frac{\rho}{1 - \rho} P(B^{res} > x), \quad x \rightarrow \infty. \end{aligned}$$

Indeed (Pollaczek-Khintchine formula for $M/G/1$):

$$E[e^{-sW}] = \frac{(1 - \rho)s}{s - \lambda + \lambda E[e^{-sB}]}$$

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SO

$$W \stackrel{d}{=} B_1^{res} + \dots + B_K^{res}.$$

Remark

$$P(B_1^{res} + \dots + B_n^{res} > x) \sim nP(B^{res} > x), \quad x \rightarrow \infty,$$

holds for subexponential (e.g., regular varying) distributions.

Crucial property: if sum is large, it is most likely due to one big term

(Catastrophe principle)

(b) Proof via LST and Bingham-Doney

If $P(B > x) \sim x^{-\nu}L(x)$, $x \rightarrow \infty$, $1 < \nu < 2$,
then (Bingham-Doney):

$$\begin{aligned}1 - E[e^{-sB^{res}}] &= 1 - \frac{1 - E[e^{-sB}]}{sEB} \\ &\sim -\frac{\Gamma(1-\nu)}{EB} s^{\nu-1} L(1/s), \quad s \downarrow 0,\end{aligned}$$

because

$$E[e^{-sB}] - 1 + sEB \sim -\Gamma(1-\nu)s^{\nu}L\left(\frac{1}{s}\right), \quad s \downarrow 0.$$

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Combine with

$$E[e^{-sW}] = \frac{1-\rho}{1-\rho E[e^{-sB^{res}}]} = \frac{1-\rho}{1-\rho + \rho(1 - E[e^{-sB^{res}}])},$$

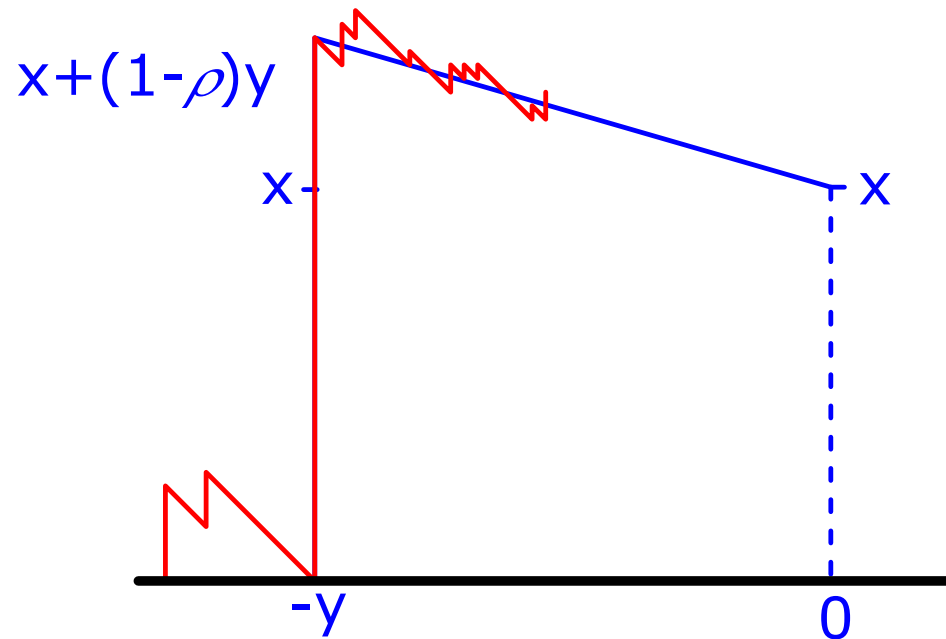
to get, for $s \downarrow 0$:

$$1 - E[e^{-sW}] \sim -\frac{\rho}{1-\rho} \frac{\Gamma(1-\nu)}{EB} s^{\nu-1} L(1/s).$$

Now again apply Bingham-Doney.

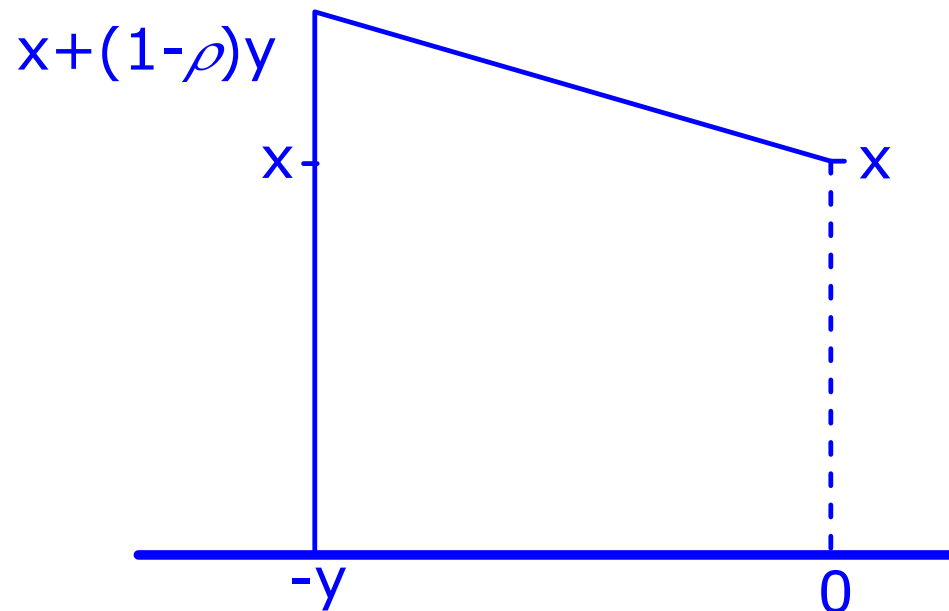
(c) Proof via sample-path argument

$$P(W > x) = P(V > x), \quad x > 0 \text{ (PASTA)}$$



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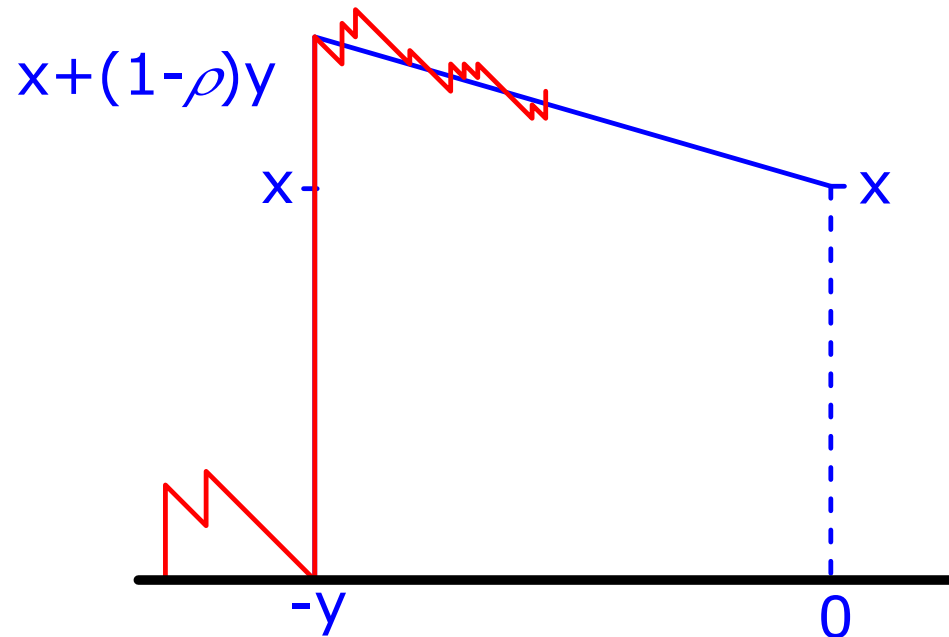
$$P(W > x) = P(V > x), \quad x > 0 \text{ (PASTA)}$$



$$P(V > x) \sim \int_{y=0}^{\infty} P(B > x + (1 - \rho)y) \lambda dy$$

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$$P(W > x) = P(V > x), \quad x > 0 \text{ (PASTA)}$$



$$\begin{aligned} P(V > x) &\sim \int_{y=0}^{\infty} P(B > x + (1 - \rho)y) \lambda dy \\ &= \int_{z=x}^{\infty} P(B > z) \lambda \frac{dz}{1 - \rho} = \frac{\rho}{1 - \rho} \int_{z=x}^{\infty} \frac{P(B > z)}{EB} dz. \end{aligned}$$

Rigorous proof: provide a lower and upper bound that asymptotically coincide.

Lower bound: “easy” .

$$P(V > x) \geq \frac{\rho}{1 - \rho + \delta} \int_{x(1+\epsilon)}^{\infty} \frac{P(B > z)}{EB} dz.$$

Use the Law of Large Numbers to show that one big jump gives this lower bound.

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Use the Law of Large Numbers to show that one big jump gives this lower bound: $A(-y, 0) > (\rho - \delta)y$ a.s.

$$P(V > x) \geq \int_{y=0}^{\infty} P(B + A(-y, 0) - y > x) \lambda dy$$

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$$\begin{aligned} P(V > x) &\geq \int_{y=0}^{\infty} P(B + A(-y, 0) - y > x) \lambda dy \\ &\geq \int_{y=0}^{\infty} P(B > x(1 + \epsilon) + y(1 - \rho + \delta)) \\ &\quad P(A(-y, 0) - \rho y \geq -\epsilon x - \delta y) \lambda dy \end{aligned}$$

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Rigorous proof: provide a lower and upper bound that asymptotically coincide.

Lower bound: “easy”.

$$P(V > x) \geq \frac{\rho}{1 - \rho + \delta} \int_{x(1+\epsilon)}^{\infty} \frac{P(B > z)}{EB} dz.$$

Upper bound: “hard”.

$$P(V > x) \leq \frac{\rho}{1 - \rho - \delta} \int_{x(1-\epsilon)}^{\infty} \frac{P(B > z)}{EB} dz + o(x^{1-\nu}).$$

Include all other scenario's (like two big jumps) and show that they can (asymptotically) be neglected.

4. $M/G/1$: busy period

De Meyer and Teugels (1980):

$$P(P > x) \sim \frac{1}{1 - \rho} P(B > (1 - \rho)x), \quad x \rightarrow \infty.$$

Proof 2: apply Bingham-Doney to $E[e^{-sP}]$.

$E[e^{-sP}]$ satisfies

$$x = E[e^{-(s+\lambda(1-x))B}].$$

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Proof 2: apply Bingham-Doney to $E[e^{-sP}]$.

$E[e^{-sP}]$ satisfies

$$x = E[e^{-(s+\lambda(1-x))B}].$$

This implies ($1 < \nu < 2$):

$$E[e^{-sP}] - 1 + s \frac{EB}{1 - \rho} \sim -\Gamma(1 - \nu)(1 - \rho)^{-(\nu+1)} s^\nu L\left(\frac{1}{s}\right), \quad s \downarrow 0.$$

Hence

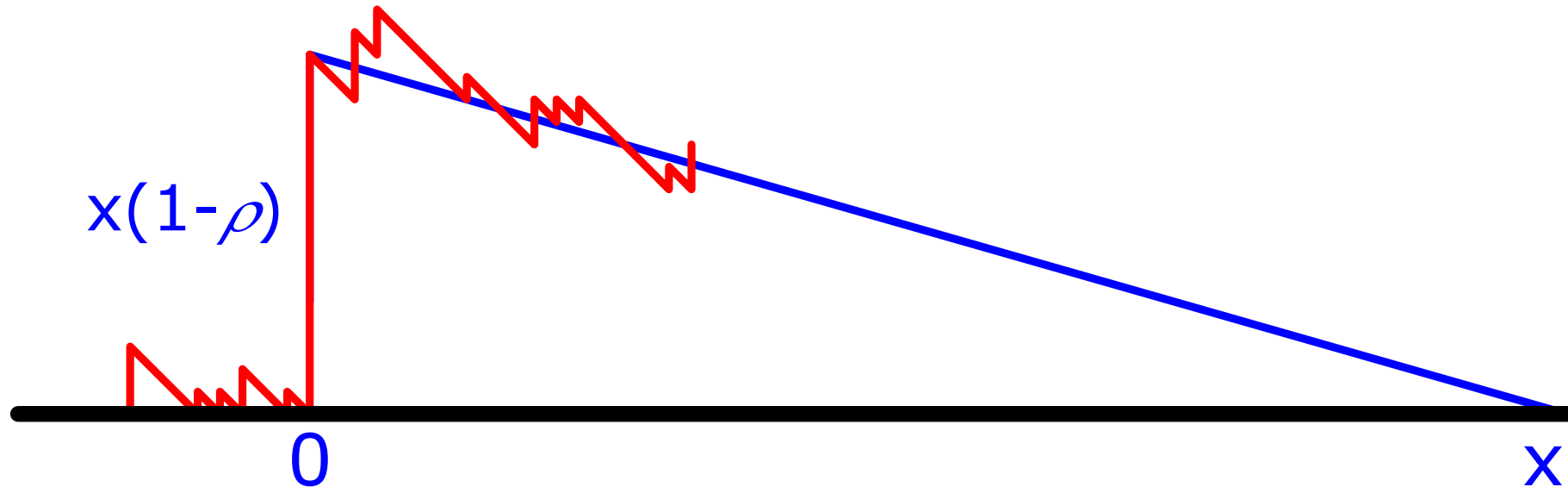
$$P(P > x) \sim \frac{1}{1 - \rho} ((1 - \rho)x)^{-\nu} L(x), \quad x \rightarrow \infty.$$

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De Meyer and Teugels (1980):

$$P(P > x) \sim \frac{1}{1 - \rho} P(B > (1 - \rho)x), \quad x \rightarrow \infty.$$

Proof 3: sample-path argument

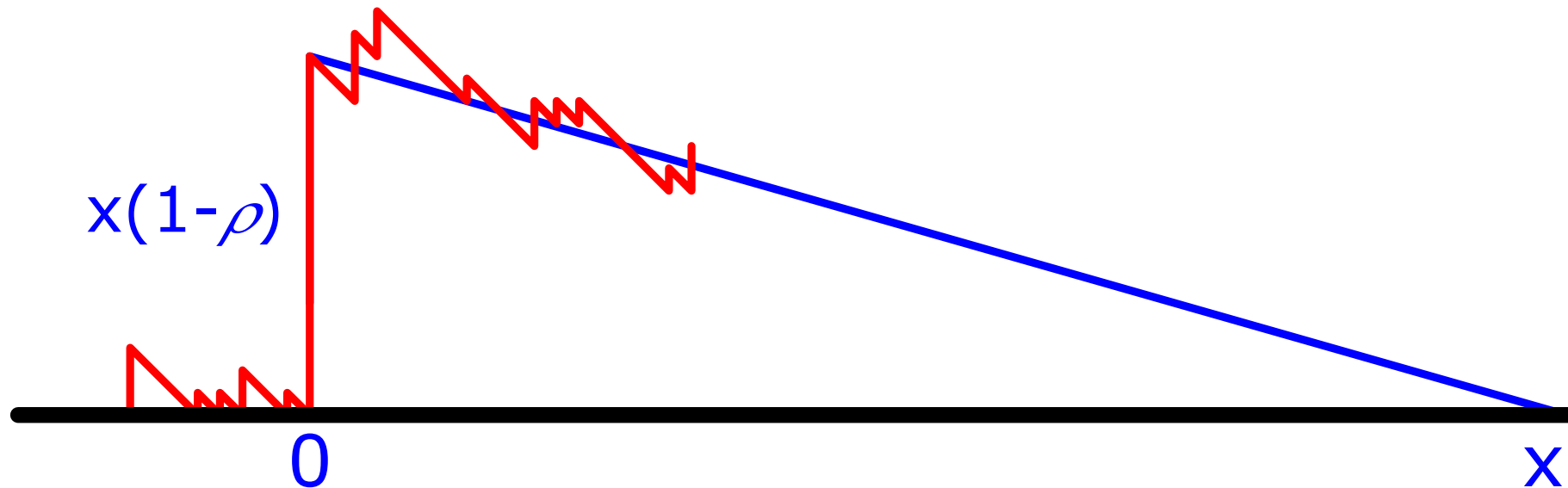


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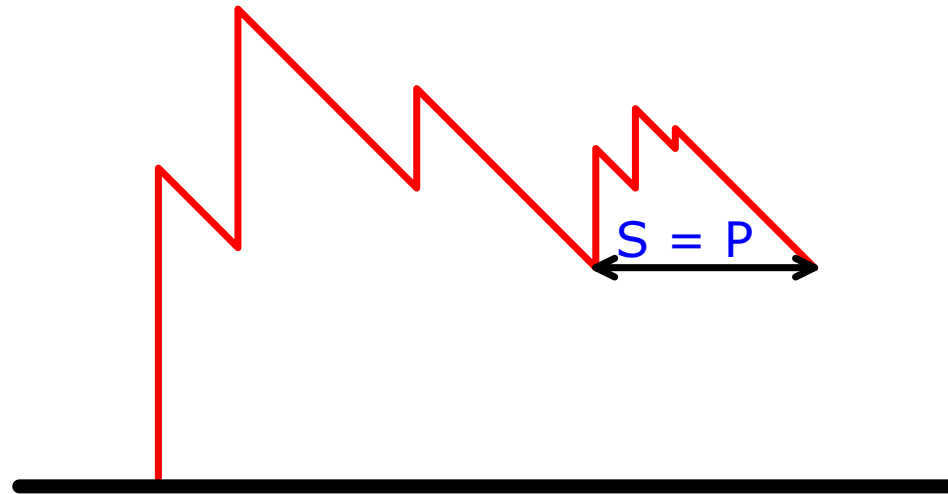
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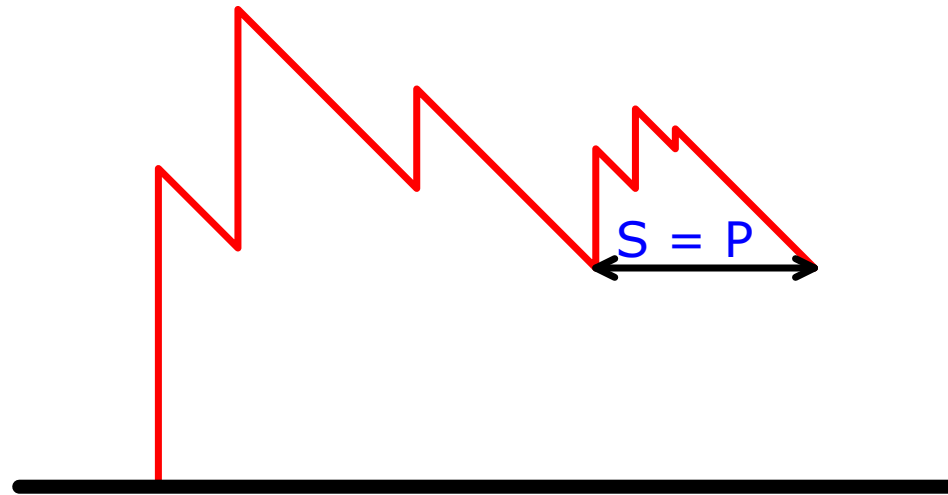


Remember that $\frac{1}{1-\rho} = E[\text{number of customers in busy period}]$.

5. $M/G/1$ LCFS preemptive resume: sojourn time



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Sojourn time $S \stackrel{d}{=} P$ busy period

Hence

$$P(S_{LCFS-PR} > x) \sim \frac{1}{1-\rho} P(B > (1-\rho)x), \quad x \rightarrow \infty.$$

Note: $O(x^{-\nu})$ instead of $O(x^{1-\nu})$!

6. Epilogue

Warning: careful with using results like

$$P(W > x) \sim \frac{\rho}{1 - \rho} P(B^{res} > x) \text{ or } Cx^{1-\nu}$$

as approximation.

Such an approximation is typically **bad**

(**Abate, Choudhury & Whitt**, QUESTA 16, 1994).

Taking a few more terms does not always help

(**Willekens & Teugels**, QUESTA 10, 1992).

Abate et al. suggest: Use their numerical algorithm for LST inversion, if such LST is known.

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Reference:

The impact of the service discipline on delay asymptotics. Sem Borst, Onno Boxma, Sindo Núñez-Queija, Bert Zwart. *Performance Evaluation* 54 (2003), 175-206.

<http://www.cwi.nl/~sindo>